# Particle-Based Approximate Inference using Random Sampling 

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## Particles

- Particles: a set of instantiations of joint distribution to all or some of the variables in the network


## Outline

- Forward Sampling
- Rejection Sampling
- Likelihood Weighting Sampling
- Importance Sampling


## Forward Sampling

- Sample the nodes in some order consistent with the partial order of the BN, so that by the time we sample a node, we have values for all its parents.





## Forward Sampling in a Bayesian network

Procedure Forward-Sample (B)
1 Let $X_{1}, \ldots, X_{n}$ be a topological ordering of X
2 For $\mathrm{i}=1, \ldots, \mathrm{n}$
$3 u_{i} \leftarrow x<P a_{x_{i}}>$ //Assignment to $P a_{X_{i}}$ in $x_{1}, \ldots, x_{i-1}$
4 Sample $x_{i}$ from $\quad P\left(X_{i} \mid u_{i}\right)$
5 return ( $x_{1}, \ldots, x_{i-1}$ )

## Absolute Error Bound

- Apply Hoeffding's bound to estimate how many samples are required to achieve an estimate whose error is bounded by $\varepsilon$, with probability at least 1- $\delta$

$$
P_{D}\left(\hat{P}_{D}(y) \notin[P(y)-\varepsilon, P(y)+\varepsilon]\right) \leq 2 e^{-2 M \varepsilon^{2}} \leq \delta
$$

Gives sample complexity:

$$
M \geq \frac{\ln (2 / \delta)}{2 \varepsilon^{2}}
$$

## Relative Error Bound

- By applying Chernoff's bound to conclude that Is also within a relative error $\varepsilon$ of the true value $P(y)$ $\hat{P}_{D}(y)$, within high probability. Specifically, we have that:

$$
P_{D}\left(\hat{P}_{D}(y) \notin P(y)(1+\in)\right) \leq 2 e^{-M P(y) \varepsilon^{2} / 3}
$$

So that:

$$
M \geq 3 \frac{\ln (2 / \delta)}{P(y) \varepsilon^{2}}
$$

## Rejection Sampling

To generate samples from $P(x \mid e)$, we can:

1. generate samples $x$ from $P(X)$,
2. reject any sample which is not compatible with e.

Problem: the number of accepted particles can be quite small. The expected number is $\mathrm{MP}(\mathrm{e})$.

- The number of samples required to achieve a low relative error grows linearly with 1/P(e)


## Likelihood Weighting

- Idea: Instead of generating samples that are rejected, simply force the samples to take on the appropriate values at observed nodes.
- Problem: particles are generated with probability that is different from $P(x)$
- Solution: each particle generated is assigned a weight that represents $P(e)$ for that sample


BBN likelihood weighting example



BBN likelihood weighting example



BBN likelihood weighting example


## BBN likelihood weighting example



BBN likelihood weighting example



BBN likelihood weighting example Second sample



BBN likelihood weighting example Second sample



## BBN likelihood weighting example

 Second sample

## Likelihood weighting

- Assume we have generated the following M samples:

- If we calculate the estimate:

$$
P(B=T \mid J=T, M=F)=\frac{\# \text { sample_with }(B=T)}{\text { \#total_sample }}
$$

a less likely sample from $P(X)$ may be generated more often.

- For example, sample than in $P(X)$
 is generated more often
- So the samples are not consistent with $P(X)$.


## Likelihood weighting

- Assume we have generated the following $M$ samples:


How to make the samples consistent?
Weight each sample by probability with which it agrees with the conditioning evidence $\mathrm{P}(\mathrm{e})$.


## Likelihood weighting

- How to compute weights for the sample?
- Assume the query $P(B=T \mid J=T, M=F)$
- Likelihood weighting:
- With every sample keep a weight with which it should count towards the estimate

$$
\begin{gathered}
\widetilde{P}(B=T \mid J=T, M=F)=\frac{\sum_{i=1}^{M} 1\left\{B^{(i)}=T\right\} w^{(i)}}{\sum_{i=1}^{M} w^{(i)}} \\
\widetilde{P}(B=T \mid J=T, M=F)=\frac{\sum_{\text {samples with } B=T \text { and } J=T, M=F} w_{B=T} w_{B=x}}{\sum_{\text {samples with any value of } B \text { and } J=T, M=F}}
\end{gathered}
$$

## Likelihood weighting

- Assume we have generated the following M samples:

- If we calculate the estimate:

$$
P(A=T \mid J=T, M=F)=\frac{\# \text { sample_with }(A=T)}{\text { \#total_sample }}
$$

a less likely sample from $P(x)$ may be generated more often. So the samples are not consistent with $\mathrm{P}(\mathrm{x})$.
How to make the samples consistent? The probability of the evidence $\mathrm{P}(\mathrm{e})$ for the sample tells us how likely the evidence is in the sample. So we can use $\mathrm{P}(\mathrm{e})$ to weight each sample and correct the bias.

## Likelihood weighting

- Assume M samples where evidence is enforced:


M
weights

- We can use $\mathrm{P}(\mathrm{e})$ to weight each sample and correct the bias.
- The correct estimate is then:

$$
\widetilde{P}(A=T \mid J=T, M=F)=\frac{\sum_{i=1}^{M} 1\left\{A^{(i)}=T\right\} w^{(i)}}{\sum_{i=1}^{M} w^{(i)}}
$$

## Likelihood weighted Particle Generation

Procedure LW-sample (B, Z=z)
//B - Bayesian network over X, Z- event in the network
1 Let $X_{1}, \ldots, X_{n}$ be a topological ordering of $X$
$2 w \leftarrow 1$
3 for $\mathrm{i}=1, \ldots, \mathrm{n}$
$4 \quad u_{i} \leftarrow x<P a_{X_{i}}>/ /$ Assignment to $P a_{X_{i}}$ in $x_{1}, \ldots, x_{i-1}$
5 If $X_{i} \notin Z$ then
6 Sample $x_{i}$ from $P\left(X_{i} \mid u_{i}\right)$
7 else
$8 \quad x_{i} \leftarrow z<X_{i}>/ /$ Assignment to $X_{i}$ in z
$9 \quad w \leftarrow w P\left(x_{i} \mid u_{i}\right)$ //Multiply weight by probability of desired value
10 return $\left(x_{1}, \ldots x_{n}\right), w$

## Likelihood Weighting

## Summary

## Likelihood Weighting

- generates M weighted particles
$<\xi[1], w[1]>, \ldots,<\xi[M], w[M]>$ using LW Sample procedure.
- Estimates the conditional probability $\mathrm{P}(\mathrm{y} \mid \mathrm{e})$ using M samples as :

$$
\hat{P}(y \mid e)=\frac{\sum_{m=1}^{M} w[m] 1\{y[m]=y\}}{\sum_{m=1}^{M} w[m]}
$$

## Importance Sampling

- Importance Sampling is a general approach for estimating the expectation of a function $\mathrm{f}(\mathrm{x})$ relative to some distribution $P(X)$ (target distribution):

$$
E_{p}[f]=\sum_{\{x\}} P(x) f(x) \quad \text { or } \quad E_{p}[f]=\int_{x} p(x) f(x) d x
$$

- Generally, we can estimate this expectation by generating samples $\mathrm{x}[1], \ldots, x[\mathrm{M}]$ from P , and then estimating

$$
\widetilde{E}_{p}[f]=\frac{1}{M} \sum_{m=1}^{M} f(x[m])
$$

## Importance Sampling

- Estimate of $\widetilde{E}_{p}[f]$ requires to sample $\mathrm{P}(\mathrm{x})$
- It might be impossible or computationally very expensive to generate samples directly from $P$.
- Because of that we might prefer to generate samples from a different distribution Q (a proposal or sampling distribution) instead
- A proposal distribution Q can be arbitrary, but it should satisfy:

$$
Q(x)>0 \text { whenever } P(x)>0
$$

## Unnormalized Importance Sampling ( P is Known)

- Since we generate samples from $Q$ instead of $P$, we must adjust our estimator to compensate for the incorrect sampling distribution.
$E_{p(X)}[f(X)]=E_{Q(x)}\left[f(x) \frac{P(x)}{Q(x)}\right]=E_{Q(x)}[f(x) w(x)]$
- We use standard estimator for expectations relative to $Q$. We generate a set of samples $D=\{x[1], \ldots, x[M]\}$ from $Q$, and estimate:

$$
\hat{E}_{D}(f)=\frac{1}{M} \sum_{m=1}^{M} f(x[m]) \frac{P(x[m])}{Q(x[m])}
$$

## Unnormalized Importance Sampling ( P is Known)

- This is an unbiased estimator: its mean for any data set is precisely the desired value
- We can estimate the distribution of this estimator around its mean: as $\mathrm{M} \rightarrow \infty$

$$
\begin{gathered}
E_{Q(X)}[f(X) w(X)]-E_{p}[f(X)] \propto N\left(0 ; \sigma_{Q}{ }^{2} / M\right) \\
w(x)=P(x) / Q(x)
\end{gathered}
$$

where

$$
\sigma_{Q}{ }^{2}=\left[E_{Q(X)}\left[(f(X) w(X))^{2}\right]\right]-\left(E_{P(X)}[f(X)]\right)^{2}
$$

## Unnormalized Importance Sampling ( P is Known)

- The variance of this estimator decreases linearly with the number of samples.
- When $f(X)=1$, the variance is simply the weighting function $P(X) / Q(X)$. Thus the more different $Q$ is from $P$, the higher the variance will be.
- The lowest variance is achieved when

$$
Q(X) \propto|f(X)| P(X)
$$

- We should avoid cases where our sampling probability $Q(X) \ll P(X) f(X)$ in any part of the space, as these cases can lead to very large or even infinite variance.


## Normalized Importance Sampling ( P is known up to a normalizing constant)

- When P is only known up to a normalizing constant $\alpha$, but we have access to a function $P^{\prime}(\mathrm{X})$, such that $P^{\prime}$ is not a normalized distribution, but $P^{\prime}(\mathrm{X})=\alpha \mathrm{P}(\mathrm{x})$
- In this context, we cannot define the weights relative to P , so

$$
\begin{aligned}
& \text { we define: } \quad w(X)=\frac{P^{\prime}(X)}{Q(X)} \\
& \begin{aligned}
E_{P(X)}[f(X)]= & \sum_{x} P(x) f(x)=\sum_{x} Q(x) f(x) \frac{P(X)}{Q(x)}=\frac{1}{\alpha} \sum_{x} Q(x) f(x) \frac{P^{\prime}(x)}{Q(x)} \\
& =\frac{1}{\alpha} E_{Q(x)}[f(X) w(X)]=\frac{E_{Q(X)}[f(X) w(X)]}{E_{Q(X)}[w(X)]}
\end{aligned}
\end{aligned}
$$

## Normalized Importance Sampling ( P is known up to a normalizing constant)

- Using an empirical estimator for both the numerator and denominator, we can estimate:

$$
\hat{E}_{D}(f)=\frac{\sum_{m=1}^{M} f(x[m]) w(x[m])}{\sum_{m=1}^{M} w(x[m])}
$$

- Although the normalized estimator is biased, its variance is typically lower than that of the unnormalized estimator. This reduction in variance often outweighs the bias term.
- Normalized estimator is often used in place of the unnormalized estimator, even in cases where $P$ is known and we can sample from it effectively.


## Proposal Distribution based on the Mutilated Belief network

Assume a Bayesian Network

- We want to calculate $P(x \mid e)$
- This is hard if we need to go opposite the links and account for the effect of evidence on nondescendants
Objective: generate particles efficiently using a simpler proposal distribution $\mathrm{Q}(\mathrm{x})$
Solution: a mutilated belief network
- Idea:
- Avoid propagation of evidence effects to nondescendants;
- Disconnect all variables in the evidence from their parents


## Mutilated Belief network

- Assume we want to calculate $\mathrm{P}(\mathrm{x} \mid \mathrm{B}=\mathrm{T}, \mathrm{J}=\mathrm{T})$ in the Alarm network
- Use $\mathrm{B}=\mathrm{T}$ and $\mathrm{J}=\mathrm{T}$ to build a mutilated network


Original network
Mutilated network

## Mutilated Belief network

- Assume the evidence is $J=j^{*}$ and $B=b^{*}$
- Original network: $P\left(E=e, A=a, M=m, J=j^{*}, B=b^{*}\right)=P\left(b^{*}\right) P(e) P\left(a \mid b^{*}, e\right) P\left(j^{*} \mid a\right) P(m \mid a)$
- Mutilated network:

$$
Q\left(E=e, A=a, M=m, J=j^{*}, B=b^{*}\right)=P(e) P\left(a \mid b^{*}, e\right) P(m \mid a)
$$

- Note that $w(x)=\frac{P(x)}{Q(x)}=P\left(b^{*}\right) P\left(j^{*} \mid a\right)$



## Mutilated Belief network

- Assume the evidence is $J=j^{*}$ and $B=b^{*}$
- Original network:

$$
P\left(E=e, A=a, M=m, J=j^{*}, B=b^{*}\right)=P\left(b^{*}\right) P(e) P\left(a \mid b^{*}, e\right) P\left(j^{*} \mid a\right) P(m \mid a)
$$

- Mutilated network:

$$
Q\left(E=e, A=a, M=m, J=j^{*}, B=b^{*}\right)=P(e) P\left(a \mid b^{*}, e\right) P(m \mid a)
$$

- Note that $w(x)=\frac{P(x)}{Q(x)}=P\left(b^{*}\right) P\left(j^{*} \mid a\right)$

So importance sampling with a proposal distribution based on mutilated network is equal to likelihood weighting


## Data-Dependent Likelihood Weighting

- Question: When to stop? How many samples do we need to see?
- Intuition: not every samples contribute equally to the quality of the estimate. A sample with high weight is more compatible with the evidence e, and may provide us with more information.
- Solution: We stop sampling when the total weight of the generated particles reaches a pre-defined value.
- Benefits: It allows early stopping in cases where we were lucky in our random choice of samples.


## Ratio Likelihood Weighting

- Estimate the conditional probability $\mathrm{P}(\mathrm{y} \mid \mathrm{e})$ in two phases: use likelihood weighing to estimate $\mathrm{P}(\mathrm{e})$ and $\mathrm{P}(\mathrm{y}, \mathrm{e})$ separately.
- Use LW M times with the argument $E=e$ to generate a set D of weighted samples ( $\xi[1], w[1]), \ldots,(\xi[M], w[M])$ use the same algorithm $M^{\prime}$ 'times with argument $Y=y, E=e$ to generate another set $D^{\prime}$ of weighted samples

$$
\left(\xi^{\prime}[1], w^{\prime}[1]\right), \ldots,\left(\xi^{\prime}[M], w^{\prime}[M]\right)
$$

- Then we can estimate:

$$
\hat{P}_{D}(y \mid e)=\frac{\hat{P}_{D^{\prime}}(y, e)}{\hat{P}_{D}(e)}=\frac{1 / M^{\prime} \sum_{m=1}^{M^{\prime}} w^{\prime}[m]}{1 / M \sum_{m=1}^{M} w[m]}
$$

## Q\&A

- Thank you!

